

0. Big Picture

Statistical Learning V
(26-11-2015)



Log(arithmetic) loss:

* my predictions are probability distributions P for Y

$$* L(Y, P) = -\log P(Y) \in [0, \infty]$$

↳ small if I gave large probability to Y

↳ large if I gave small probability to Y .

Maximum Likelihood = ERM for log loss

$$\begin{aligned} \text{ML: } \arg\max_{\beta} \prod_{i=1}^N P_{\beta}(Y_i | X_i) &= \arg\min_{\beta} \sum_{i=1}^N -\log P_{\beta}(Y_i | X_i) \\ &= \arg\min_{\beta} \sum_{i=1}^N L(Y_i, P_{\beta}(\cdot | X_i)) \end{aligned}$$

↑
loss when learning is log loss
in big picture.

Statistical Learning I

(26-11-2015)

(1)

0. Big Picture

1. Surrogate losses

2. Logistic regression

3. Discriminative vs generative methods

~~4. Discriminative vs generative methods~~ Unsupervised learning:

5. Clustering: K-means, EM with Gaussian mixtures

1. Surrogate losses

Classifications: find \hat{f} with small expected 0/1-loss

$$EPE(\hat{f}) = \mathbb{E}_{x,y} [L(y, \hat{f}(x))]$$

0 if $\hat{f}(x)$ predicts y correctly
1 if $\hat{f}(x)$ predicts y incorrectly

Empirical Risk Minimization:

$$\hat{f} = \underset{f \in \mathcal{F}}{\operatorname{argmin}} \frac{1}{N} \sum_{i=1}^N L(y_i, f(x_i))$$

Suppose \mathcal{F} set of all linear functions: $\{f(x) = x^T \beta \mid \beta \in \mathbb{R}^p\}$

$$y \in \{-1, +1\} \quad L(y, f(x)) = \begin{cases} 0 & \text{if } \operatorname{sign}(f(x)) = y \\ 1 & \text{if } \operatorname{sign}(f(x)) \neq y \end{cases}$$

Difficulties:

- ERM solution not unique for linearly separable data
- computational problems for non-separable data

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Surrogate/proxy loss L' :

Care about loss L , but train classifier with loss L'

Want: L large $\Leftrightarrow L'$ large

L small $\Leftrightarrow L'$ small

Usually: surrogate L' convex in parameters, so can optimize efficiently

Two most important surrogates: logistic loss, hinge loss

\uparrow
today

\uparrow
next week

Predict correctly iff $\text{sign}(x^T \hat{\beta}) = y \Rightarrow y \cdot x^T \hat{\beta}$ larger is better

(see Figure of surrogate losses)

2. Logistic Regression

(How can we use the power of linear regression for classification?)

Assume 2 classes for simplicity.

Bad idea: $\hat{P}(y=1|x) = x^T \hat{\beta}$ ← least squares for classification

E.g. y_i is going to the beach on day i .

x_{i3} is temperature in $^{\circ}\text{C}$

$$\hat{\beta}_3 = \frac{1}{5}$$

Then $+10^{\circ}\text{C} \Rightarrow x^T \hat{\beta} + 2 \Rightarrow \hat{P}(y=1|x) + 2 > 1 \quad \nabla$

Probabilities do not behave linearly.

Better: $\log P_{\beta}(Y=1|X) = x^T \beta \iff P_{\beta}(Y=1|X) = e^{x^T \beta}$

$x^T \hat{\beta} + 2 \implies P_{\hat{\beta}}(Y=1|X)$ increases by factor e^2 ,
but can still go outside $[0, 1]$

Logistic Regression:

log odds is linear

$$\log \frac{P_{\beta}(Y=1|X)}{P_{\beta}(Y=-1|X)} = x^T \beta \iff \begin{cases} P_{\beta}(Y=1|X) = \frac{e^{x^T \beta}}{1 + e^{x^T \beta}} = \frac{1}{1 + e^{-x^T \beta}} \\ P_{\beta}(Y=-1|X) = \frac{1}{1 + e^{x^T \beta}} \end{cases}$$

$$P_{\beta}(Y=1|X) = \frac{1}{1 + e^{-x^T \beta}}$$

$P_{\hat{\beta}}(Y=1|X) > \frac{1}{2} \iff x^T \hat{\beta} > 0 \implies$ decision boundary:
 x s.t. $x^T \hat{\beta} = 0$
 is linear

logistic loss = log loss for logistic regression:

$$L(Y|\hat{\beta}) = -\log P_{\hat{\beta}}(Y|X) = \log(1 + e^{-Y X^T \hat{\beta}})$$

Train with maximum likelihood = ERM for logistic loss

$$\underset{\beta}{\text{minimize}} \sum_{i=1}^N \log(1 + e^{-y_i x_i^T \beta})$$

No formula for minimum, but can find it by convex optimization

Important extensions:

$$\text{minimize} \sum_{i=1}^N \log(1 + e^{-y_i x_i^T \beta}) + \lambda \text{pen}(\beta)$$

L_2 -penalty: $\text{pen}(\beta) = \sum_{j=1}^p \beta_j^2$ (like ridge regression)

L_1 -penalty: $\text{pen}(\beta) = \sum_{j=1}^p |\beta_j|$ (like lasso)

Considerations for penalization similar as for squared error.

3. Discriminative vs Generative Models

Probabilistic model: $\mathcal{P} = \{P_\beta(x, y) \mid \beta \in \mathbb{R}^p\}$

Generative

(imagine data generated by first choosing class y , then features x given y .)

e.g. spam

estimate $P(x, y)$

LDA, Naive Bayes, ...

overkill ~~for~~ for classification

Discriminative

estimate $P(y|x)$

Logistic regression, ...

less overkill, closer to class boundary, which is only thing we are interested in

$$\arg\min_{\beta} \sum_{i=1}^N -\log P_{\beta}(x_i, y_i)$$

$$= \arg\min_{\beta} \sum_{i=1}^N -\log P_{\beta}(y_i | x_i) + \sum_{i=1}^N -\log P_{\beta}(x_i)$$

acts as regularizer, so reduces variance \Rightarrow better for small N

$$\arg\min_{\beta} \sum_{i=1}^N -\log P_{\beta}(y_i | x_i)$$

$$= \arg\min_{\beta} \sum_{i=1}^N -\log P_{\beta}(y_i | x_i) + \sum_{i=1}^N -\log P^*(x_i)$$

some does not depend

\nearrow true distribution of x

better fit of true distribution for large N if ~~model~~ $P^* \notin \mathcal{P}$

Two methods are generative-discriminative pair if one estimates $P_{\beta}(x, y)$ and other $P_{\beta}(y|x)$ in same model \mathcal{P} .

E.g.: ~~logistic~~ LDA - logistic regression for continuous features x_j
 Naive Bayes - logistic regression for binary features

Logistic Regression vs LDA

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$x_{ij} \in \mathbb{R}$ (continuous features)

LR: $\log \frac{\hat{P}(y=1|x)}{\hat{P}(y=-1|x)} = x^T \hat{\beta} + \hat{\beta}_0 \iff \hat{P}(y=1|x) = \frac{e^{x^T \hat{\beta} + \hat{\beta}_0}}{1 + e^{x^T \hat{\beta} + \hat{\beta}_0}}$ ← writing intercept separately

LDA: ^{linear} decision boundary = $\{x : \hat{P}(y=1|x) = \hat{P}(y=-1|x)\}$
 $= \{x : \log \frac{\hat{P}(y=1|x)}{\hat{P}(y=-1|x)} = 0\}$

in fact, $\log \frac{\hat{P}(y=1|x)}{\hat{P}(y=-1|x)} = x^T \hat{\alpha} + \hat{\alpha}_0$ is always linear in x ,
not just for x s.t. it is 0.

parameters $\hat{\alpha}, \hat{\alpha}_0$ are
a function of LDA parameters
 $\hat{\mu}_k, \hat{\Sigma}, \hat{\pi}_k$ ← possible to choose s.t. $\hat{\alpha} = \hat{\beta}$
but LDA does not do that. $\hat{\alpha}_0 = \hat{\beta}_0$

so $\hat{P}(y=1|x) = \frac{e^{x^T \hat{\alpha} + \hat{\alpha}_0}}{1 + e^{x^T \hat{\alpha} + \hat{\alpha}_0}}$ ← has same form as
for logistic regression.

LDA and LR are generative-discriminative pair

↳ expect LDA to be better for smaller N ,
LR for larger N . (unless LDA model is correct)

Efron, 1975: if LDA model correct, then LR will perform
as well as LDA if LR has 30% more data,
so cannot be a lot worse.

Logistic Regression vs Naive Bayes

$x_{ij} \in \{0, 1\}$ (binary features), ~~needs not to do~~

$j = 1, \dots, p$ ~~array coding feature~~

\hookrightarrow transform into $x_i = \begin{pmatrix} x_{i1} \\ \vdots \\ x_{ip} \end{pmatrix}$ where x_{ij} is nr of a in i -th feature vector x_i

Will show that

(*) $\log \frac{\hat{P}(y=1|X)}{\hat{P}(y=-1|X)} = x^T \hat{\alpha} + \hat{\alpha}_0$ for naive Bayes (NB)

determined by NB parameters for multinomial models and by $\hat{\theta}_{ik}^k$

Hence NB - LR form generative - discriminative pair.

So expect NB better for small N , LR better for large N .

NIPS

Ng, Jordan, 2001 make this precise

* Asymptotic performance as $N \rightarrow \infty$: LR always better than NB

* But LR needs $N \geq O(p)$ to reach asymptotic performance

whereas NB needs $N \geq O(\log p)$ to reach asymptotic performance under some technical assumptions

To show (*): $\log \frac{\hat{P}(y=1|X)}{\hat{P}(y=-1|X)} = \log \frac{\hat{P}(X|y=1)\hat{P}(y=1)}{\hat{P}(X|y=-1)\hat{P}(y=-1)} = \log \frac{f_1(X)\pi_1}{f_{-1}(X)\pi_{-1}}$

f_k : multinomial model on binary features = Bernoulli model
 simplify parametrization: $\hat{\theta}_k$ is prob $x_{ij}=1$ for class k
 $1-\hat{\theta}_k$ is prob $x_{ij}=0$

~~Handwritten derivations and scribbles at the bottom of the page, including terms like $\log \frac{\hat{\theta}_k^{x_{ij}} (1-\hat{\theta}_k)^{1-x_{ij}}}{\hat{\theta}_k^{x_{ij}} (1-\hat{\theta}_k)^{1-x_{ij}}}$~~

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$$\log \frac{\hat{f}_2(x) \hat{\pi}_1}{\hat{f}_1(x) \hat{\pi}_2} = \log \frac{\hat{\theta}_1^{n_1} (1 - \hat{\theta}_1)^{n_0}}{\hat{\theta}_2^{n_1} (1 - \hat{\theta}_2)^{n_0}} + \log \frac{\hat{\pi}_1}{\hat{\pi}_2}$$

$$= n_0 \cdot \log \frac{1 - \hat{\theta}_1}{1 - \hat{\theta}_2} + n_1 \cdot \log \frac{\hat{\theta}_1}{\hat{\theta}_2} + \log \frac{\hat{\pi}_1}{\hat{\pi}_2}$$

$$= x^T \hat{\alpha} + \hat{\alpha}_0 \quad \text{for } \hat{\alpha} = \begin{pmatrix} \log \frac{1 - \hat{\theta}_1}{1 - \hat{\theta}_2} \\ \log \frac{\hat{\theta}_1}{\hat{\theta}_2} \end{pmatrix} \quad \hat{\alpha}_0 = \log \frac{\hat{\pi}_1}{\hat{\pi}_2}$$

possible to choose $\hat{\theta}_1, \hat{\theta}_2, \hat{\pi}_1, \hat{\pi}_2$ s.t. $\hat{\alpha} = \beta$
 $\hat{\alpha}_0 = \beta_0$,

but NB does not do that.

Unsupervised Learning

4.3 Intro

Supervised: $T = \left(\begin{matrix} y_1 \\ x_1 \end{matrix} \right), \dots, \left(\begin{matrix} y_N \\ x_N \end{matrix} \right)$ ← can evaluate every method by EPE

Unsupervised: $T = \{x_1, \dots, x_N\}$

What can we do?

Clustering: split data into groups of points (clusters) that are similar.

- ① What do you mean by "similar"?
- ② What kind of clusters are you looking for?

K-means

① Squared Euclidean distance between $x_i, x_{i'}$:

$$d(x_i, x_{i'}) = \sum_{j=1}^p (x_{ij} - x_{i'j})^2$$

② Find K clusters s.t. distance to mean within each cluster is small.

$c(i) \in \{1, \dots, K\}$ is cluster assigned to x_i .

$$\text{minimize } \sum_{k=1}^K \sum_{i: c(i)=k} d(x_i, \mu_k) \quad (*)$$

where μ_k = mean of points in cluster k

NB. Mistake in book! (multiplies by N_k in each cluster)

Algorithm:

1. Initialize cluster assignment C
2. Set μ_k to current mean in cluster k
3. Given means μ_1, \dots, μ_K , assign each x_i to nearest mean to get new C
4. repeat from 2 until no changes in C .

- Converges to local minimum

- ~~often~~ ^{may} choose K s.t. increasing K does not reduce $(*)$ very much.

Gaussian Mixtures and EM

- like k-means, but with "soft" cluster assignments

- for simplicity: explain for two clusters

each cluster is a Gaussian modelled by

cluster 1: $X \sim \mathcal{N}(\mu_1, \sigma_1^2)$ w.p. $1-\pi$

cluster 2: $X \sim \mathcal{N}(\mu_2, \sigma_2^2)$ w.p. π

probability density: $(1-\pi)\phi_{\mu_1, \sigma_1^2}(x) + \pi\phi_{\mu_2, \sigma_2^2}(x)$

To find parameters:

ML is difficult numerically

+ gives bad solution with $\hat{\sigma}_1 = 0$

$\hat{\mu}_1 = x_i$ for some i .

Solution: algorithm very similar to K-means

Hidden variables $\Delta_i = \begin{cases} 0 & \text{if } x_i \text{ in cluster 1} \\ 1 & \text{if } x_i \text{ in cluster 2} \end{cases}$

1. Initializes parameters $\hat{\pi}, \hat{\mu}_1, \hat{\sigma}_1, \hat{\mu}_2, \hat{\sigma}_2$ ($\hat{\sigma}_1 > 0.5, \hat{\sigma}_2 > 0.5$)

2. Soft cluster assignment:

$$\hat{\gamma}_i = E[\Delta_i | \hat{\theta}, T] = \frac{\hat{\pi} \phi_{\hat{\mu}_2, \hat{\sigma}_2}(x_i)}{\hat{\pi} \phi_{\hat{\mu}_2, \hat{\sigma}_2}(x_i) + (1 - \hat{\pi}) \phi_{\hat{\mu}_1, \hat{\sigma}_1}(x_i)}$$

3. Update parameter estimates for clusters:

$$\hat{\mu}_1 = \frac{\sum_{i=1}^N (1 - \hat{\gamma}_i) x_i}{\sum_i (1 - \hat{\gamma}_i)} \quad \hat{\mu}_2 = \frac{\sum_{i=1}^N \hat{\gamma}_i x_i}{\sum_i \hat{\gamma}_i}$$

$$\hat{\sigma}_1^2 = \frac{\sum_{i=1}^N (1 - \hat{\gamma}_i) (x_i - \hat{\mu}_1)^2}{\sum_i (1 - \hat{\gamma}_i)} \quad \hat{\sigma}_2^2 = \frac{\sum_{i=1}^N \hat{\gamma}_i (x_i - \hat{\mu}_2)^2}{\sum_i \hat{\gamma}_i}$$

$$\hat{\pi} = \frac{\sum_{i=1}^N \hat{\gamma}_i}{\sum_{i=1}^N \hat{\gamma}_i + \sum_{i=1}^N (1 - \hat{\gamma}_i)} = \frac{\sum_{i=1}^N \hat{\gamma}_i}{N}$$

4. repeat from 2 until convergence

- converges to local minimum.